

# Model Selection:

The Red-haired stepchild of Bayesian  
Analysis

# Why model selection?

- Inference in Bayes (like in ML) is conditional on the model being true
- We don't know the model is true
  - Affects reliability of posterior distribution, precision, etc

# Frequentist/ Classical approaches

- Likelihood ratio testing
- AIC

# Bayes- the dog's lunch

- Bayes Factors'
- Information criteria
  - AIC
  - BIC
  - DIC
- Cross validation
- Reversible jump MCMC

# Bayes statement of problem

The problem till now:

$$[\underline{\theta}, X] = [\underline{\theta} | X][X] = [X | \underline{\theta}][\underline{\theta}]$$

$$[\underline{\theta} | X] \propto [X | \underline{\theta}][\underline{\theta}]$$

likelihood      prior

Averaging by simulation to get individual parameters

# With multiple models--


$$[\{\underline{\theta}_m\}, m, X] = [\{\underline{\theta}_m\}, m | X][X] = [X | \{\underline{\theta}_m\}][\{\underline{\theta}_m\} | m][m]$$

$$[\{\underline{\theta}_m\}, m | X] \propto [X | \{\underline{\theta}_m\}][\{\underline{\theta}_m\} | m][m]$$

Likelihood for  
Model m



Priors  
For model m



Prior that model m is true



# What we want...

Posterior inference on parameters, averaged across models

$$[\underline{\theta} | X] \propto \int_m [X | \{\underline{\theta}_m\}][\{\underline{\theta}_m\} | m][m]dm$$

Posterior probability that model  $m$  is true

$$[m | X] \propto \int_{\theta} [X | \{\underline{\theta}_m\}][\{\underline{\theta}_m\} | m][m]d\underline{\theta}_m$$

Consider 2 models..

$$[m_1 | X] \propto [X | m_1][m_1]$$

$$[m_2 | X] \propto [X | m_2][m_2]$$

Posterior odds                      Bayes factor                      Prior odds

$$\frac{[m_1 | X]}{[m_2 | X]} = \frac{[X | m_1]}{[X | m_2]} \times \frac{[m_1]}{[m_2]}$$

# Bayes factor/ aka “likelihood ratio”

$$BF = \frac{[X | m_1]}{[X | m_2]}$$

$$\log(BF) = \log([X | m_1]) - \log([X | m_2])$$

$\gg 1$ : evidence for model 1 ( $\log BF \gg 0$ )

$\ll 1$ : evidence for model 2 ( $\log BF \ll 0$ )

$\sim 1$ : evidence for either (neither) model ( $\log BF \sim 0$ )

>2 models

$$[m_i | X] = \frac{[X | m_i][m_i]}{\sum_{k=1}^m [X | m_k][m_k]}$$

# Bayesian Information Criterion ( $\Delta BIC$ )

- “Penalty” for number of parameters
- Related to Bayes Factor
- Similar to  $\Delta AIC$

$$\Delta BIC = -2 \ln(BF) - \underbrace{(p_1 - p_2) \log(n)}_{\substack{\uparrow \\ (\rightarrow 0 \text{ as } n \rightarrow \infty)}}$$

Asymptotically = BF

# AIC

$$AIC_i = -2 \ln([X | m_i]) + 2 p_i$$

Deviance for  
model i



Complexity penalty



$$BIC_i = -2 \ln([X | m_i]) + p_i \log(n)$$

# Implementation in Winbugs

- Deviance can be monitored as a node or calculated in model

$$Deviance_i = -2\ln([X | m_i])$$

So:

$$\Delta BIC = Deviance_1 - Deviance_2 - (p_1 - p_2)\log(n)$$

$$BF = \exp[Deviance_1 - Deviance_2] \times 0.5)$$

$$AIC_i = Deviance_i + 2p_i$$

# Winbugs example

- [binomial\\_models.odc](#)

# DIC

- Proposed by Spiegelhalter et al
- Estimates “effective number of parameters”
- Enabled by DIC ‘tool’ in WINBUGS

$$DIC_i = \overline{Deviance_i} + p_D$$

$$p_D = \overline{Deviance_i} - Deviance(\bar{\theta})$$

Mean of posterior deviances (averaged over parameters, data)– the usual

Deviance evaluated at posterior parameter means

# Issues

- BF ok for posterior prediction of data (see [DIC notes](#))
- AIC counts parameters ok in non-hierarchical models but not in hierarchical ones
- BIC can overstate model strength as  $n$  gets big
- DIC
  - Sometimes doesn't work without intervention
  - Can result in negative (nonsensical)  $p_D$

# Alternative: RJMCMC

- Reversible Jump MCMC
- Treats model probability as another sampling problem
  - Jumps between parameter spaces of alternative models
  - Time spent in model  $i$  used to calculate marginal probability of model

# RJ-MCMC

Want to sample from

$$[\{\underline{\theta}_m\}, m | X] \propto [X | \{\underline{\theta}_m\}][\{\underline{\theta}_m\} | m][m]$$

Essentially we need a rule for jumping back and forth between models  $j$  and  $k$

$$[\{\underline{\theta}_j\}, j | X] \Leftrightarrow [\{\underline{\theta}_k\}, k | X]$$

Jumping distribution  $J(k, j)$

Acceptance probability

$$\alpha(j, k) = \min \left[ 1, \frac{[\{\underline{\theta}_k\}, k | X] J(k, j)}{[\{\underline{\theta}_j\}, j | X] J(j, k)} \right]$$

Proceeds iteratively in a similar manner to Metropolis-Hastings

# RJ-MCMC

- Provides a fully Bayesian treatment of model probability using MCMC
- In theory should work for any identifiable model
- As yet no general algorithm
- Potentially very slow to converge

# Information theoretic distance measures (Jim Peterson)

## Akaike information criteria

$$\text{AIC} = -2\ln(\text{likelihood}) + 2*K$$

- *Assumes no true model*
- *Tends to overfit models*

## Bayesian or Schwartz's information criteria

$$\text{BIC} = -2\ln(\text{likelihood}) + K\ln N$$

- *Assumes true model is contained in the set*
- *Tends to overstate evidence with large datasets*

## Consistent Akaike information criteria

$$\text{CAIC} = -2\ln(\text{likelihood}) + K\ln(N+1)$$

- *Similar to BIC*

## Deviance information criteria

DIC = used in MCMC modeling (more later)

**RJ-MCMC**

# Assignment

- For your examples or other data construct 2-3 alternative models (increasing or decreasing complexity but same data)
- Compare models using BF, BIC, AIC, and DIC
- Make a recommendation based on model selection criteria and/or GOF